

Recap:

$$H(X|Y) \leq H(X)$$

$$H(X,Y) \leq H(X) + H(Y)$$

Mutual information

$$I(X;Y) = H(X) - H(X|Y) = H(Y) - H(Y|X)$$

KL divergence

$$D(p\parallel q) = \sum_{i=1}^m p_i \log \frac{p_i}{q_i}$$

$$I(X;Y) = D(p(x,y) \parallel p(x) \cdot p(y))$$


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Consider a channel with input distribution  $p = (p_1, \dots, p_m)$

$$p \rightarrow \left[ \overline{p(y_j|x_i)} \right] \rightarrow \pi$$

$X \sim$  input distr.      channel      output distr.  $\sim \pi$

Let matrix  $W = (w_{ij})_{\substack{i,j=1,\dots,m \\ j=1,\dots,d}} \in \mathbb{R}^{m \times d}$ ,

$W$  is a stochastic matrix, i.e.,  $\sum_{j=1}^d w_{ij} = 1$ ,  $\forall i = 1, \dots, m$ .

Output distribution  $\pi = (\pi_1, \dots, \pi_d)$  is obtained as

$$\pi = p W$$

Lemma 2.1.13. For any distr.  $p, q$  with support  $X = \{x_1, \dots, x_m\}$  and stochastic matrix  $\underline{w}$

$$W = (p(y_j|x_i))_{ij} \in \mathbb{R}^{m \times d},$$

$$D(p||q) \geq D(pW||qW). \quad \blacksquare$$

Proof. Use the log-sum inequality La. 2.1.7

$$\sum_i a_i \log \frac{a_i}{b_i} \geq \sum_i a_i \log \frac{\sum_j a_j}{\sum_j b_j}.$$

$$\begin{aligned} D(p||q) &= \sum_{i=1}^m p(x_i) \log \frac{p(x_i)}{q(x_i)} \\ &= \sum_{i=1}^m \sum_{j=1}^d \underbrace{p(x_i) p(y_j|x_i)}_{a_i} \log \frac{p(x_i) p(y_j|x_i)}{q(x_i) p(y_j|x_i)} \\ &\geq \sum_{j=1}^d p \underline{w}_j \log \frac{p \underline{w}_j}{q \underline{w}_j} \\ &= D(pW||qW). \quad \blacksquare \end{aligned}$$

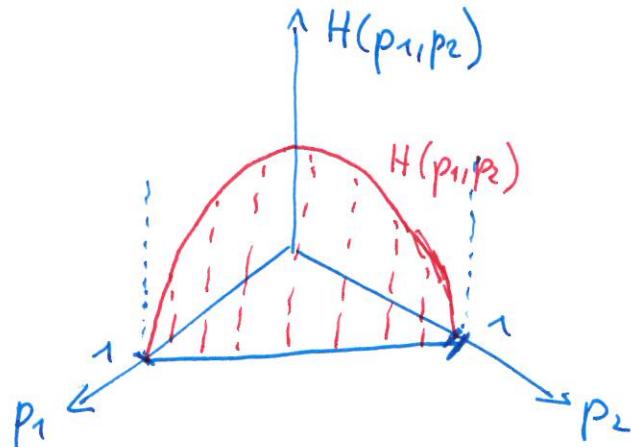
$\underbrace{q_i}_{b_i}$   
 $\underbrace{\frac{p(x_i) p(y_j|x_i)}{q(x_i) p(y_j|x_i)}}_{b_i}$   
 $w = (\underline{w}_1, \dots, \underline{w}_d)$

Theorem 2.1.14.

$H(p)$  is a concave function of  $p = (p_1, \dots, p_m)$ .

Figure w.r.t. Th. 2.1.14

2-dim.  $p = (p_1, p_2)$ ,  $p_1, p_2 \geq 0$ ,  $p_1 + p_2 = 1$



Proof. Let  $u = (\frac{1}{m}, \dots, \frac{1}{m})$  be uniform distribution.

$$D(p \parallel u) = \sum_{i=1}^m p_i \log \frac{p_i}{\frac{1}{m}} = \log m - H(p)$$

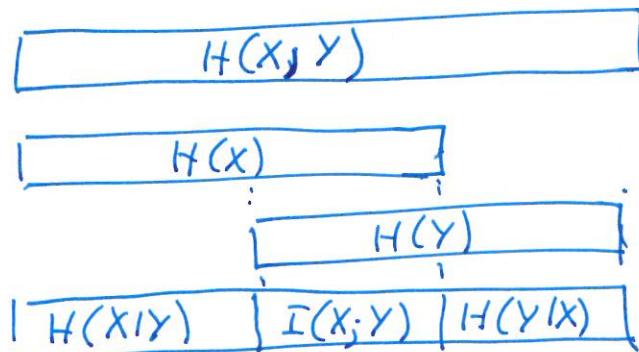
Hence  $H(p) = \log m - D(p \parallel u)$

By Th. 2.1.12 b)

$$\begin{aligned} & D(\lambda p + (1-\lambda)q \parallel \lambda u + (1-\lambda)u) \\ & \leq \lambda D(p \parallel u) + (1-\lambda)D(q \parallel u), \text{ i.e.,} \\ & \text{convexity in the first argument.} \end{aligned}$$

Thus,  $H(p)$  is a concave fct. of  $p$ .  $\blacksquare$

Collecting the quantities in a single picture.



Connections are easy to derive, e.g.-,

$$I(X; Y) = H(X) - H(X|Y)$$

$$H(Y|X) = H(X, Y) - H(X)$$

$$I(X; Y) = H(X) + H(Y) - H(X, Y)$$

## 2.2. Inequalities

Def. 2.2.1. Random variables  $X, Y, Z$  are said to have the Markovian property if the joint p.m.f. satisfies

$$p(x, y, z) = p(x) p(y|x) p(z|y).$$

For ~~the~~ Notation:  $X \rightarrow Y \rightarrow Z$ .

For  $X \rightarrow Y \rightarrow Z$  the conditional distn. of  $Z$  depends only on  $Y$  and is conditionally independent of  $X$ .

Lemma 2.2.2.

a) If  $X \rightarrow Y \rightarrow Z$ , then

$$p(x, z | y) = p(x|y) \cdot p(z|y).$$

b) If  $X \rightarrow Y \rightarrow Z$ , then  $Z \rightarrow Y \rightarrow X$ .

c) If  $Z = f(Y)$ , then  $X \rightarrow Y \rightarrow Z$ .  $\perp$

Proof.

$$\begin{aligned} a) \quad p(x, z | y) &= \frac{p(x, y, z)}{p(y)} = \frac{p(x) p(y|x) p(z|y)}{p(y)} \\ &= \frac{p(x|y) p(z|y)}{p(y)} = p(x|y) p(z|y). \end{aligned}$$

$$\begin{aligned} b) \quad p(x, y, z) &= p(x) p(y|x) p(z|y) \\ &= p(x|y) \frac{p(z|y)}{p(y)} \frac{p(z)}{p(z)} \\ &= p(z) p(y|z) p(x|y), \text{ i.e., } Z \rightarrow Y \rightarrow X. \end{aligned}$$

c)  $Z = f(Y)$ , then

$$p(x, y, z) = \begin{cases} p(x, y), & \text{if } z = f(y) \\ 0, & \text{otherwise} \end{cases}$$

Hence  $p(x, y, z) = p(x, y) \mathbb{1}(z = f(y)) = p(x) p(y|x) p(z|y)$ ,

$$\text{since } p(z|y) = \begin{cases} 1, & z = f(y) \\ 0, & \text{otherwise} \end{cases}$$



Theorem 2.2.3. (Data-processing inequality)

If  $X \rightarrow Y \rightarrow Z$ , then  $I(X;Z) \leq \min\{I(X;Y), I(Y;Z)\}$

'No processing of  $Y$  can increase the information that  $Y$  contains about  $X$ '

Proof By the chain rule

$$\begin{aligned} I(X;Y,Z) &= I(X;Z) + \overbrace{I(X;Y|Z)}^{\geq 0} \\ &= I(X;Y) + \overbrace{I(X;Z|Y)}^{=0} \end{aligned}$$

Since  $X$  and  $Z$  are cond. independent given  $Y$ :

$I(X;Z|Y)=0$ . Since  $I(X;Y|Z) \geq 0$ , we have

$$I(X;Z) \leq I(X;Y)$$

Equality holds iff  $I(X;Y|Z)=0$ , i.e.,  $X \rightarrow Z \rightarrow Y$ .

$I(X;Y) \leq I(Y;Z)$  is shown analogously.  $\blacksquare$

Assume  $X, Y$  r.v. with support  $\mathcal{X} = \{x_1, \dots, x_u\}$ .

Define  $p_e = P(X \neq Y)$ , the 'error probability'.

Theorem 2.2.4. (Fano inequality)

$$H(X|Y) \leq H(p_e) + p_e \log(u-1).$$

This implies that  $p_e \geq \frac{H(X|Y) - \log 2}{\log(u-1)}$   $\blacksquare$

$$H(p_e) = -p_e \log p_e - (1-p_e) \log(1-p_e)$$

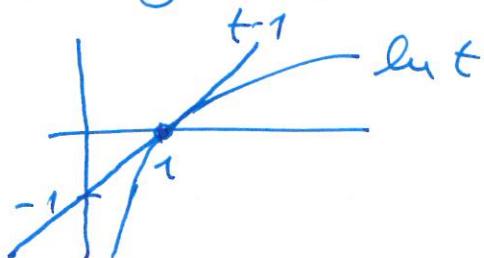
Beweis: Proof.

$$(i) \quad H(X|Y) = \sum_{x \neq y} p(x,y) \log \frac{1}{p(x|y)} + \sum_x p(x,x) \log \frac{1}{p(x|x)}$$

$$(ii) \quad p_e \log(m-1) = \sum_{x \neq y} \log(m-1) p(x,y)$$

$$(iii) \quad H(p_e) = -p_e \log p_e - (1-p_e) \log (1-p_e)$$

$$(iv) \quad \ln t \leq t-1, t \geq 0$$



$$H(X|Y) = p_e \log(m-1) - H(p_e)$$

$$\stackrel{(i),(ii)}{=} \sum_{x \neq y} p(x,y) \log \frac{p_e}{p(x|y)(m-1)}$$

$$+ \sum_x p(x,x) \log \frac{(1-p_e)}{p(x|x)}$$

$$\leq (\log e) \left[ \sum_{x \neq y} p(x,y) \left( \frac{p_e}{p(x|y)(m-1)} - 1 \right) \right.$$

$$\left. + \sum_x p(x,x) \left( \frac{1-p_e}{p(x|x)} - 1 \right) \right]$$

$$= (\log e) \left[ \frac{p_e}{m-1} \sum_{x \neq y} p(y) - \sum_{x \neq y} p(x,y) \right.$$

$$\left. + (1-p_e) \sum_x p(x) - \sum_x p(x,x) \right]$$

$$= (\log e) [p_e - p_e + (1-p_e) - (1-p_e)] = 0$$

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