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$$X \Rightarrow H(X) = - \sum_{i=1}^m p_i \cdot \log p_i$$

$$\text{Prob}(X=x_i) = p_i, \quad i=1 \dots m$$

Theorem 2.1.B.

$$a) \quad 0 \stackrel{(i)}{\leq} H(X) \stackrel{(ii)}{\leq} \log m$$

"=" in (i) $\Leftrightarrow X$ has a singleton dist., i.e.,
 $\exists k_i : \text{Prob}(X=k_i) = 1$

"=" in (ii) $\Leftrightarrow X$ is uniformly dist., i.e.,

$$\text{Prob}(X=k_i) = \frac{1}{m} \quad \forall i=1, \dots, m$$

$$b) \quad 0 \stackrel{(i)}{\leq} H(X|Y) \stackrel{(ii)}{\leq} H(X)$$

"=" in (i) $\Leftrightarrow \text{Prob}(X=k_i | Y=y_j) = 1 \quad \forall i,j$
 with $\text{Prob}(X=k_i, Y=y_j) > 0$, i.e.,
 X is totally dependent on Y .

"=" (iii) $\Leftrightarrow X, Y$ are stoch. independent.

$$c) \quad H(X) \stackrel{(i)}{\leq} H(X, Y) \stackrel{(ii)}{\leq} H(X) + H(Y)$$

"=" in (i) $\Leftrightarrow Y$ is totally dependent on X

"=" in (ii) $\Leftrightarrow X, Y$ are stoch. independent.

$$d) \quad H(X|Y, Z) \leq \min \{ H(X|Y), H(X|Z) \}$$

Proof:

$$\begin{aligned}
 \text{a) (ii)} \quad H(K) &= -\sum_{i=1}^m p_i \cdot \log p_i = \sum_{i=1}^m p_i \cdot \log \frac{1}{p_i} \\
 &\stackrel{\text{Lemma 2.1.6}}{\leq} \log \left(\sum_{i=1}^m p_i \cdot \frac{1}{p_i} \right) \\
 &= \log \left(\sum_{i=1}^m 1 \right) = \log m
 \end{aligned}$$

(c) obvious

b) Exercise

c) (i) By the chain rule, Th. 2.1.5.:

$$\begin{aligned}
 H(X, Y) &= H(X) + \underbrace{H(Y|X)}_{\substack{\text{equality follows from } \stackrel{>0}{\text{b) (ii)}}}} \geq H(X) \\
 (\text{iii}) \quad 0 &\stackrel{(\text{b})}{\leq} H(X) - H(X|Y) \stackrel{\text{Th. 2.1.5}}{=} H(X) - [H(X, Y) - H(Y)] \\
 &\Rightarrow H(X) + H(Y) \geq H(X, Y) \\
 &\text{with equality from b) (i).}
 \end{aligned}$$

d) Analogous to b) (ii'). ■

Ref. Cover & Thomas, *Math.*

Definition 2.19.

Let X, Y, Z discrete r.v.

$$\begin{aligned} I(X; Y) &= H(X) - H(X|Y) \\ &= H(Y) - H(Y|X) \end{aligned}$$

is called mutual information or synergropy of X and Y .

$$I(X; Y | Z) = H(X|Z) - H(X|Y, Z)$$

is called conditional mutual information of X and Y given Z .

Interpretation: $I(X; Y)$ is the reduction in uncertainty about X when Y is given or the amount of information about X provided by Y .

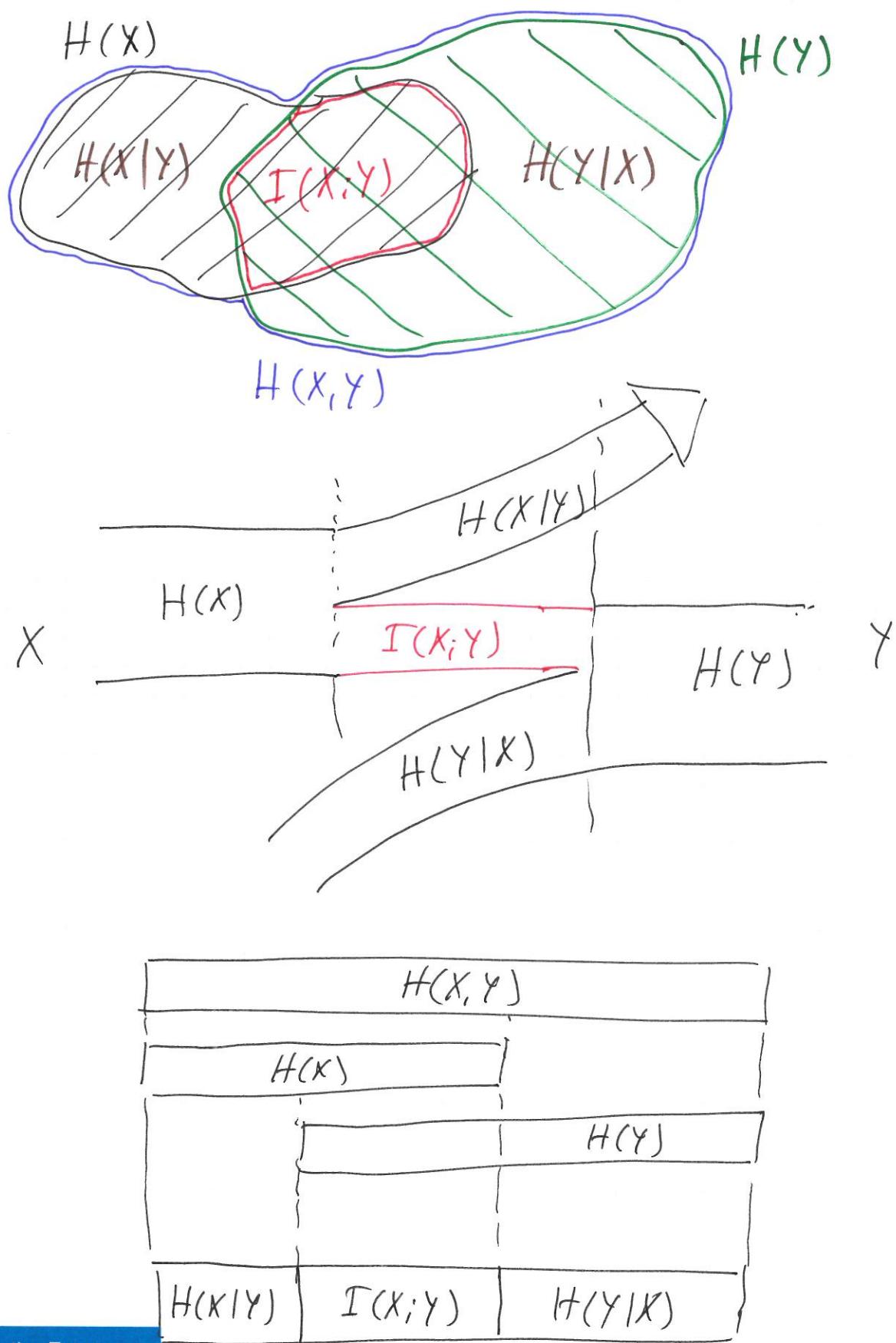
see. Figures on the next page.

$$I(X; Y) = H(X) - H(X|Y)$$

$$H(Y|X) = H(X, Y) - H(X)$$

$$I(X; Y) = H(X) + H(Y) - H(X, Y)$$

4.



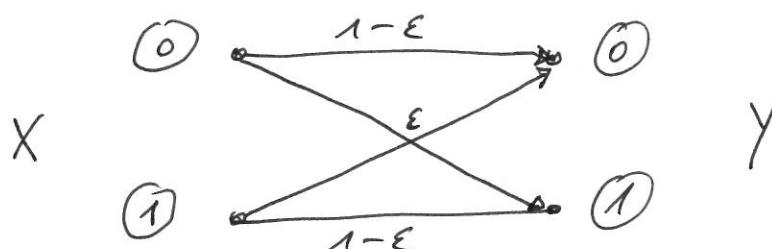
Note: by Thm 2.1.8 b) $I(X; Y) \geq 0$.

By definition it holds that:

$$\begin{aligned}
 I(X; Y) &= - \underbrace{\sum_i p(x_i) \cdot \log p(x_i)}_{= H(X)} + \sum_{ij} p(x_i, y_j) \cdot \log p(x_i | y_j) \\
 &= - \sum_{ij} p(x_i, y_j) \cdot \log p(x_i) + \sum_{ij} p(x_i, y_j) \cdot \log p(x_i | y_j) \\
 &= \sum_{ij} p(x_i, y_j) \cdot \log \frac{p(x_i | y_j)}{p(x_i)} \\
 &= \sum_{ij} p(x_i, y_j) \cdot \log \frac{p(x_i, y_j)}{p(x_i) \cdot p(y_j)} \quad (*)
 \end{aligned}$$

which shows symmetry in X and Y .

Example 2.1.10. Binary symmetric channel (BSC)



symbol error with probability ε , $0 \leq \varepsilon \leq 1$.

Hence:

$$P(Y=0 | X=0) = P(Y=1 | X=1) = 1-\varepsilon$$

$$P(Y=0 | X=1) = P(Y=1 | X=0) = \varepsilon$$

Assume that input symbols are uniformly distributed: $P(X=0) = P(X=1) = \frac{1}{2}$.

Then for the joint distributions:

$$P(X=0, Y=0) = P(Y=0 | X=0) \cdot P(X=0) = (1-\varepsilon) \cdot \frac{1}{2}$$

;

		$X \backslash Y$	0	1	
		0	$\frac{1}{2}(1-\varepsilon)$	$\frac{\varepsilon}{2}$	$\frac{1}{2}$
		1	$\frac{\varepsilon}{2}$	$\frac{1}{2}(1-\varepsilon)$	$\frac{1}{2}$
			$\frac{1}{2}$	$\frac{1}{2}$	

Further: $P(X=0 | Y=0) = \frac{P(X=0, Y=0)}{P(Y=0)} = 1-\varepsilon$

$$P(X=1 | Y=1) = 1-\varepsilon$$

$$P(X=0 | Y=1) = P(X=1 | Y=0) = \varepsilon$$

For $\log = \log_2$

$$\Rightarrow H(K) = H(Y) = -\frac{1}{2} \cdot \log \frac{1}{2} - \frac{1}{2} \cdot \log \frac{1}{2} = 1 \text{ bit.}$$

$$H(K, Y) = 1 - (1-\varepsilon) \cdot \log(1-\varepsilon) - \varepsilon \cdot \log \varepsilon$$

$$H(K|Y) = -(1-\varepsilon) \cdot \log(1-\varepsilon) - \varepsilon \cdot \log \varepsilon$$

$$0 \leq I(K; Y) = 1 + (1-\varepsilon) \cdot \log(1-\varepsilon) + \varepsilon \cdot \log \varepsilon \leq 1$$

Def. 2.1.11 \rightarrow (Kullback-Leibler divergence)

Let $p = (p_1, \dots, p_m)$, $q = (q_1, \dots, q_m)$ be stoch. vectors.

$$D(p \parallel q) = \sum_{i=1}^m p_i \cdot \log \frac{p_i}{q_i}$$

is called KL divergence between p and q
 (or relative entropy).

$D(p \parallel q)$ measures the divergence (distance, dissimilarity) between distributions p and q .
 However, it is not a metric, neither symmetric nor satisfies the triangle inequality.

It measures how difficult it is for p to pretend it were q .

Theorem 2.1.12.

- a) $D(p \parallel q) \geq 0$ with " $=$ " iff $p = q$.
- b) $D(p \parallel q)$ is convex in the pair (p, q) .
- c) $I(X; Y) = D\left(\left(p(x_i, y_j)\right)_{ij} \parallel \left(p(x_i) \cdot p(y_j)\right)_{ij}\right)$

proof.

- a) By definition and Cor. 2.1.8.
- b) use the log-sum inequality Lemma 2.1.7.
Let p, r , and q, s be stoch. vectors.
For all $i = 1, \dots, m$ it holds that

$$\begin{aligned} & (\lambda p_i + (1-\lambda) \cdot r_i) \cdot \log \frac{\lambda p_i + (1-\lambda) \cdot r_i}{\lambda q_i + (1-\lambda) \cdot s_i} \\ & \leq \lambda p_i \log \frac{\lambda p_i}{\lambda q_i} + (1-\lambda) \cdot r_i \log \frac{(1-\lambda) \cdot r_i}{(1-\lambda) \cdot s_i} \end{aligned}$$

Summing over all $i = 1, \dots, m$, it follows $\forall \lambda \in [0, 1]$

$$D(\lambda p + (1-\lambda) r \parallel \lambda q + (1-\lambda) s) \leq \lambda D(p \parallel q) + (1-\lambda) D(r \parallel s)$$

- c) By definition. □

Note: $D(p \parallel q) \neq D(q \parallel p)$